

# EPSE 592: Design & Analysis of Experiments

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# Last Time

- Bernoulli, Binomial, Normal r.v.s
- Sample statistics
- Standard errors

# Last Time

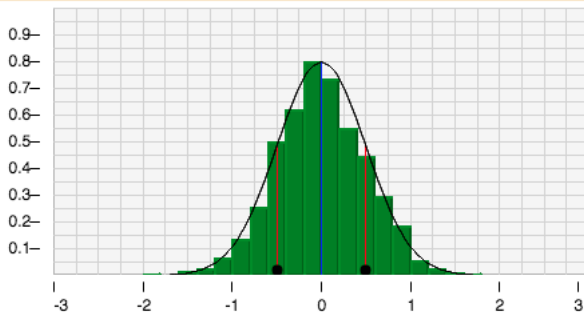
- Confidence intervals
- Central Limit Theorem
- Hypothesis testing
- Test statistics and p-values
- $Z$ -test,  $t$ -test,  $F$ -test

# Sample Statistics

- In practice, we study a random variable by observing its values on only a *sample*.
- Studying this sample allows us to infer properties of the actual random variable if the sample is random and representative.
- This is basically what applied statistics is all about!

# Sample Statistics

- We can approximate a r.v.'s PMF or PDF by plotting a *histogram* of our sample data.



Standard Deviation = 0.5

Visit: <http://www.shodor.org/interactivate/activities/NormalDistribution/>

# Sample Statistics

- We can get a sense of the “typical” value of our r.v. by calculating a *sample mean*, *sample median*, or *sample mode*.
- Let  $\{X_1, \dots, X_n\}$  denote a random sample of  $n$  independent observations from the random variable  $X$ . We define the **sample mean** by:

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i.$$

- sample median = 50th percentile of sample data
- sample mode = most commonly observed value in sample data
- Remember: these can all be different!

- We can get a sense of the spread or dispersion (variability) of our r.v. by calculating a *sample variance*.
- Let  $\{X_1, \dots, X_n\}$  denote a random sample of  $n$  independent observations from the random variable  $X$ . We define the **sample variance** by:

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2.$$

# Sample Statistics vs. Properties of Random Variables

- Although the definitions of *expectation* and *sample mean*, and of *variance* and *sample variance*, look very similar, they are fundamentally different.
  - Sample mean and variance are *functions of the data/sample*. Different samples will generate different values for sample mean/variance *even if the samples are from the same population*.
  - Expectations and variances of random variables are idealized quantities. They are inherent properties of the random phenomenon we are studying. We usually cannot calculate them in practice; we can only estimate them via our *sample* approximations.



# Standard Errors

- Because sample statistics are random (i.e. not fixed) quantities, they are genuine random variables on their own!
- Thus, they have expectations, variances, std. devs. of their own.
- Terminology: the **standard error** of a sample statistic is simply its standard deviation.
- If  $T$  denotes a sample statistic, then we usually write  $SE(T)$  to denote its standard error.
- In practice, standard errors are functions of the sample size and the original variability in the population from which we sampled our data.

# Confidence Intervals

- A confidence interval is a way of summarizing a sample statistic (e.g. sample mean) and its standard error at once.
- An (approximate) 95% confidence interval for the expectation (population mean),  $\mu_X$ , of a continuous random variable  $X$  from a random sample  $\{X_1, \dots, X_n\}$ , for large  $n$ , is

$$[\bar{X} - 2 \cdot SE(\bar{X}), \bar{X} + 2 \cdot SE(\bar{X})]$$

- Notice, this CI depends on the sample; i.e., it is a statistic.
- **Interpretation:** if we resample 100 times and calculate the 95% confidence interval for each new sample, then approximately 95 of those CIs will contain the true (unknown) population mean.

# Central Limit Theorem

## Central Limit Theorem (CLT)

Let  $\{X_1, \dots, X_n\}$  denote a random sample of  $n$  independent observations from a common distribution with finite mean  $\mu$  and finite variance  $\sigma^2$ . Recall the sample mean is given by

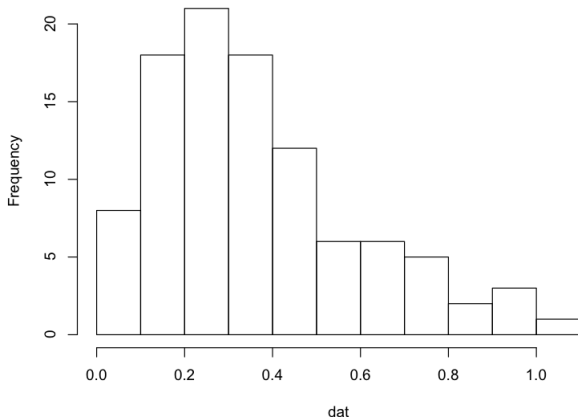
$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i.$$

Then, for  $n$  large,  $\bar{X}$  is approximately distributed as  $N(\mu, \sigma^2/n)$ .

- This is one of the most important theorems of classical statistics. Tells us all about how the sample mean behaves for an independent random sample from *any* common distribution with finite mean and variance.

# Central Limit Theorem: example

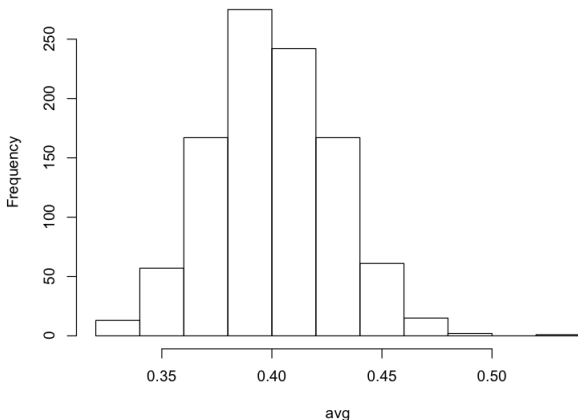
Histogram of **random sample** of size 100 from a very skewed (Gamma) random variable.



Sample mean is 0.366 for this particular set of 100 sample data points.

# Central Limit Theorem: example continued

Histogram of the **sample means** of 1000 random samples (each of size 100) from the same very skewed (Gamma) random variable.



Notice the histogram looks quite Normal! (CLT at work)

# Central Limit Theorem

- **Moral:** CLT allows us to treat the **sample mean** of *any* random phenomenon as a normal random variable, *as long as our sample size is big enough*.
- This will allow us to assign a measure of uncertainty to our sample mean estimate, e.g. by constructing *confidence intervals*.
- For small sample sizes, either the random phenomenon itself must follow a normal distribution, or we need to use other (nonparametric) statistical methods.

# Statistical Hypothesis Testing

- Nearly all quantitative science is based around the idea of stating and testing quantifiable hypotheses about study objects of interest.
- Point Null Hypothesis Testing (PNHT) is the most common option in virtually all applied disciplines.

# Statistical Hypothesis Testing

Basic recipe of PNHT:

- (1) Identify parameter of interest.
- (2) Define null hypothesis,  $H_0$ , of *no effect*.
- (3) Define a *test statistic*  $T$  (a function of the data) such that the larger  $T$  is, the less consistent our data are with  $H_0$ .
- (4) Collect data and then compute test statistic:  $t_{obs}$ .
- (5) Compute p-value =  $\Pr(|T| \geq t_{obs} \mid H_0)$ ; if p-value small enough, then conclude data are **inconsistent** with  $H_0$ .

Example:

- (1) Difference in mean response between treatment groups  $X$  and  $Y$
- (2)  $H_0 : \mu_X = \mu_Y$
- (3)  $T =$  standardized difference in sample means
- (4) Collect data; compute  $t_{obs} = (\bar{X} - \bar{Y})/SE$
- (5) Calculating p-value requires knowing distribution of  $T$  given  $H_0$ ....

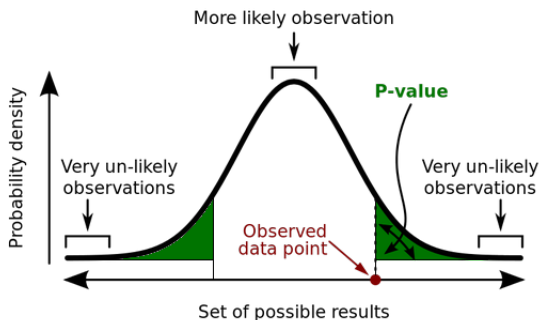


# A Closer Look at P-values

- Formally, we define

$$p\text{-value} = \Pr(|T| \geq t_{obs} \mid H_0), \text{ usually.}$$

- Interpretation:** the p-value is the probability of observing a test statistic as or more extreme than the one observed for our sample, given that the null hypothesis is true.



# A Closer Look at P-values

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$$\text{p-value} = \Pr(|T| \geq t_{obs} \mid H_0), \text{ usually.}$$

- Interpretation:** the p-value is the probability of observing a test statistic as or more extreme than the one observed for our sample, given that the null hypothesis is true.
- So a big p-value means the observed test statistic is “typical” under  $H_0$ . Therefore, the data are consistent with  $H_0$ .
- A small p-value means the observed test statistic is *not* “typical” under  $H_0$ . Therefore, the data are inconsistent with  $H_0$ .

# A Closer Look at P-values

- Formally, we define

$$\text{p-value} = \Pr(|T| \geq t_{obs} \mid H_0), \text{ usually.}$$

- Recall definition of conditional probability:

$$\Pr(|T| \geq t_{obs} \mid H_0 \text{ true}) = \frac{\Pr(|T| \geq t_{obs}, \text{ and } H_0 \text{ true})}{\Pr(H_0 \text{ true})}.$$

- With this in mind, how could the p-value be small?

# Z-test for Difference of Means

## Proposition

If  $X$  and  $Y$  data come from **normal distributions** with the **same known variance**  $\sigma^2$  but possibly different means, then the test statistic

$$T = (\bar{X} - \bar{Y})/SE$$

also follows a normal distribution, with mean  $\mu_X - \mu_Y$  and variance  $\sigma^2/n$ , where  $n$  denotes the sample size. Therefore, we can calculate

$$p\text{-value} = Pr(|T| \geq t_{obs} \mid H_0)$$

since  $T$  is  $N(0, \sigma^2/n)$  under  $H_0$ .

Notice that we do *not* assume anything about  $\mu_X$  and  $\mu_Y$ , the quantities we are trying to study. Assuming  $H_0$  (i.e. hypothesis of no difference) allows us to bypass any quantitative assumptions on these parameters.

# T-test for Difference of Means

In practice, we are never going to actually know the value of  $\sigma^2$ . Instead, we can estimate it by the *sample variance*. This will allow us to *estimate* the SE.

## Proposition

If  $X$  and  $Y$  are  $n$  data points coming from **normal distributions** with the **same (unknown) variance**  $\sigma^2$  but possibly different means, then the test statistic

$$T = (\bar{X} - \bar{Y}) / \widehat{SE}$$

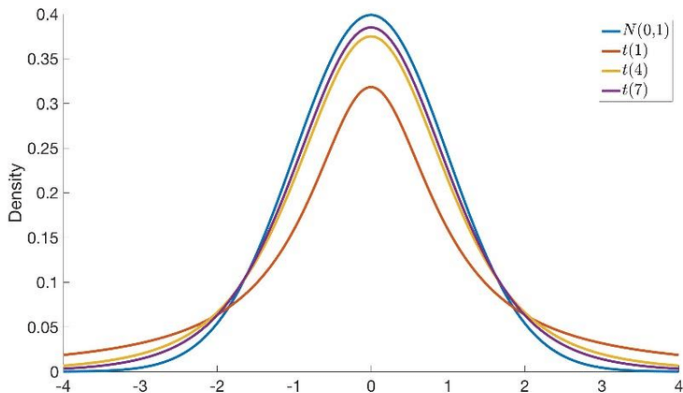
follows a Student- $t$  distribution on  $(n - 2)$  degrees of freedom, with mean  $\mu_X - \mu_Y$ . Therefore, we can calculate

$$p\text{-value} = \Pr(|T| \geq t_{obs} \mid H_0)$$

since  $T$  is  $t_{n-2}$  (a known probability distribution) under  $H_0$ .

# Student- $t$ Random Variables

- Student- $t$  random variables look like normal distributions, but with *heavy tails*; i.e. extreme events are more likely.



## Example: t-test (independent samples)

- Suppose we have annual gross income figures (in \$1000's) for a random sample of 10 British Columbians and 10 Albertans:

BC	44	45	46	34	48	42	68	44	52	51
AB	59	50	83	43	65	70	67	77	52	51

- Can use an *independent samples t-test* to test the null hypothesis

$$H_0 : \mu_{BC} = \mu_{AB}.$$

Here, we assume that the BC subjects were sampled independently from the AB subjects.

- Must also check assumptions of t-test:
  - (1) independence of observations
  - (2) normality of data
  - (3) homogeneity of variances (homoskedasticity)

## Example: t-test (independent samples) in Jamovi

- Enter data as two columns (income, province) in Jamovi

8	44	BC
9	52	BC
10	51	BC
11	59	AB
12	50	AB
13	83	AB
14	43	AB

- Click “Analyses” tab, then “T-Tests”, then “Independent Samples T-Test”
- Assign “Income” to dependent variable
- Assign “Province” to grouping variable
- Test statistic, degrees of freedom of (theoretical) Student-*t* random variable, and p-value will appear in output on right side of screen
- Click on appropriate boxes to produce tests/plots for assumptions, confidence intervals, etc.



## Example: t-test (paired samples)

- Suppose we have test scores for 9 first-year calculus students before and after taking a weekend review workshop on pre-calculus topics (algebra, geometry, trigonometry).

before	77	78	82	67	75	91	53	66	70
after	75	80	90	70	70	90	65	74	77

- Can use a *paired samples t-test* to test the null hypothesis

$$H_0 : \mu_{\text{before}} = \mu_{\text{after}}.$$

Here, we are measuring the *same subjects* at two different time points; thus, their responses are **dependent**. A paired t-test accounts for this lack of independence.

- Must also check assumptions of this t-test:
  - (1) normality of data

## Example: t-test (paired samples) in Jamovi

- Enter data as two columns (before, after) in Jamovi

	before	after
1	77	75
2	78	80
3	82	90
4	67	70
5	75	70

- Click “Analyses” tab, then “T-Tests”, then “Paired Samples T-Test”
- Assign “before” and “after” to paired variables
- Test statistic, degrees of freedom of (theoretical) Student-*t* random variable, and p-value will appear in output on right side of screen
- Click on appropriate boxes to produce tests/plots for assumptions, confidence intervals, etc.

# Testing for Equality of Variances

- Equality of variances is an assumption for an *unpaired t*-test.
- But how can we rigorously test if two variances are (statistically) equal?

Sample 1	51	53	49	40	55	56	49	48	42	51
Sample 2	47	45	35	50	70	62	49	37	57	63

- Can calculate sample variances of the two samples: use formula or use “Descriptives” tab in Jamovi.
- $S_1 = 5.15$  and  $S_2 = 11.4$
- But are these statistically different? Remember: sample variances are *random variables*. So is this observed difference in sample variances meaningful, given the inherent randomness of the data?

# F-test for Inequality of Variances

## Proposition

Suppose we draw  $n_1$  sample points from the random variable  $X$  and  $n_2$  sample points from the random variable  $Y$ . If these  $X$  and  $Y$  data come from **normal distributions** with possibly different means and possibly different variances  $\sigma_1^2$  and  $\sigma_2^2$ , then the test statistic

$$T = \frac{S_1^2}{S_2^2}$$

follows a Fisher-F distribution on  $(n_1 - 1)$  numerator degrees of freedom and  $(n_2 - 1)$  denominator degrees of freedom under the null hypothesis

$$H_0 : \sigma_1^2 = \sigma_2^2.$$

As before, small p-value should reflect when  $T$  is an “extreme” value under  $H_0$ . This happens if  $S_1 \gg S_2$  or if  $S_1 \ll S_2$ .

# Testing for Equality of Variances

- Back to our example:

Sample 1	51	53	49	40	55	56	49	48	42	51
Sample 2	47	45	35	50	70	62	49	37	57	63

- $S_1 = 5.15$  and  $S_2 = 11.4$
- In Jamovi, follow the procedure for an independent samples t-test from before.
- Under “Assumption Checks,” click the box for “Equality of variances.”
- Produces Levene's Test, (essentially) the test statistic  $S_1^2/S_2^2$  compared against its theoretical  $F$  distribution under  $H_0$ .

# F-Tests

- $F$ -tests always take the form of a ratio of variances.
- When the two variances describe normal data, then the ratio of sample variances is a Fisher- $F$  random variable.
- Will rely heavily on this all term: we will usually assume model errors are normally distributed. So can use  $F$ -tests to compare if the variance of one model is significantly less than another model (i.e. if one model explains more of the variation in the data than another model).

## Summary of statistical tests so far...

- $Z$ -test for testing difference of two group means from (approx.) normal data with known variance.
- $T$ -test for testing difference of two group means from (approx.) normal data with unknown variance. Paired and unpaired versions.
- $F$ -test for testing difference of two group variances from (approx.) normal data.

Note: the CLT implies that we can use *all* these tests for non-normal data as long as we have large enough sample sizes.

## Summary of statistical tests so far...

- $Z$ -test for testing difference of **two** group means from (approx.) normal data with known variance.
- $T$ -test for testing difference of **two** group means from (approx.) normal data with unknown variance. Paired and unpaired versions.
- $F$ -test for testing difference of **two** group variances from (approx.) normal data.

Note: the CLT implies that we can use *all* these tests for non-normal data as long as we have large enough sample sizes.

- What about when we want to test for a difference between *more than two* group means?



## Example: three experimental groups of interest

Suppose we are interested in studying how amount of higher education correlates with self-reported anxiety levels. We have a survey designed to measure anxiety and give it to 18 people at UBC: 6 who have obtained Bachelor's degrees, 6 who have obtained Master's degrees, and 6 who have obtained PhDs (chosen how?).

Bachelor's	Master's	PhD
6.2	6.2	6.9
5.8	6.9	9.0
6.0	6.2	7.7
5.9	7.7	9.1
6.6	6.8	8.3
6.2	7.9	8.0

**Table:** Self-reported anxiety levels, 10 point scale. 18 respondents.

# Example: three experimental groups of interest

- Could perform 3 independent-samples t-tests to test the 3 null hypotheses:

- $H_{0,1} : \mu_B = \mu_M$

Independent Samples T-Test

		statistic	df	p
A	Student's t	-2.63	10.0	0.025

- $H_{0,2} : \mu_M = \mu_P$

Independent Samples T-Test

		statistic	df	p
C	Student's t	2.71	10.0	0.022

- $H_{0,3} : \mu_B = \mu_P$

Independent Samples T-Test

		statistic	df	p
E	Student's t	-5.73	10.0	<.001

## Example: three experimental groups of interest

- Could perform 3 independent-samples t-tests to test the 3 null hypotheses:
  - $H_{0,1} : \mu_B = \mu_M \implies p\text{-value} < 0.05$
  - $H_{0,2} : \mu_M = \mu_P \implies p\text{-value} < 0.05$
  - $H_{0,3} : \mu_B = \mu_P \implies p\text{-value} \ll 0.05$
- But what about inflated Type I error?

# Type I and Type II Errors

- Recall: when p-value small, conclude data inconsistent with  $H_0$ .
- Recall: when p-value large, conclude data consistent with  $H_0$ .
- Whenever we make a decision about a hypothesis based on a p-value, we have a chance of making an error.

	$H_0$ true	$H_0$ false
data inconsistent with $H_0$	Type I error <i>false positive</i>	Correct decision <i>true positive</i>
data consistent with $H_0$	Correct decision <i>true negative</i>	Type II error <i>false negative</i>

# Type I and Type II Errors

- Traditionally, we set a predetermined *significance level*,  $\alpha$ , such that

$$\Pr(\text{Type I error}) = \Pr(p - \text{value} < \alpha \mid H_0 \text{ true}) = \alpha.$$

- Then  $\alpha$ , sample size, variability, and choice of test determine

$$\Pr(\text{Type II error}) = \Pr(p - \text{value} > \alpha \mid H_0 \text{ false}) = \beta.$$

- The *confidence level*, or *specificity*, of a test is defined as

$$\Pr(p - \text{value} > \alpha \mid H_0 \text{ true}) = 1 - \alpha.$$

- The *power*, or *sensitivity*, of a test is defined as

$$\Pr(p - \text{value} < \alpha \mid H_0 \text{ false}) = 1 - \beta.$$

# Type I and Type II Errors

- In practice,  $\alpha = 0.05$  is a common choice.
- Note: all of  $1 - \alpha$ ,  $\beta$ , and  $1 - \beta$  are determined once  $\alpha$  has been fixed, the data have been collected, and the choice of analysis made.
- Good studies will strive to have  $1 - \beta \geq 0.80$ . Most studies will have much lower power.

	Given $H_0$ true	Given $H_0$ false
Pr(data inconsistent with $H_0$   ...)	$\alpha$	$(1 - \beta)$
Pr(data consistent with $H_0$   ...)	$(1 - \alpha)$	$\beta$

# Multiple Testing

- Each time we conduct a statistical test of hypothesis, we have a chance of committing a Type I or Type II error.
- The choice of  $\alpha$  controls our chance of Type I error for a single test.
- Thus, if our study requires more than one test, each one has a chance of error.
- Thus, if our study requires more than one test, we should be concerned with the *family-wise* error rate: the probability of committing *at least one* Type I error.